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# Import Price Misalignment after the Crisis: A New Keynesian Perspective

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### Motivations

- Will exchange rate adjustments be sufficient for rebalancing U.S. external deficits? Maybe not, because of **low exchange-rate passthrough** to import prices.
- How are U.S. import prices related to the production cost of its trading partners, especially Asian countries? If there are any **misalignments**, how to explain them?
- An essential source of persistent trade imbalances might be associated with **import-price stickiness**.
- Then how to **identify the role of price stickiness** abstracting from that of monetary policy?
- Are region-, country-, or industry-specific factors responsible for misalignments?

### Passthrough to Import Prices



#### **Import Price Misalignment**

Import Price Index (All Goods But Petroleum)

Exchange Rate Adjusted, Trade Weighted Index of Trading Partner Domestic Manufacturing PPI's

• Following Marquez and Thomas (2006)

US Import Price Misalignment (2000=1)



Figure 2a US Import Price Misalignment by Global Region of Origin (2004= 1)



Figure 2b: Import Price Misalignment: Asian Subregions (2004 = 1)



## State of Play

- U.S. imports have historically been cheap relative to foreign production costs.
  - U.S. dollar has declined but not reflected in more expensive imports.
- The import price misalignment measure declines most extremely among Pacific Rim countries, especially ASEAN.

## Import Price Passthrough

- Well documented: slow pass-through and sticky import prices.
- Need optimizing model to:
  - Control for endogenous costs: Exchange rates vs. marginal costs.
  - Immunize from the Lucas critique: Slow price changes may be attributable to monetary policy response.
- Asian monetary policy challenges for "Full Information" model.

### Strategy

- Estimate the degree of price stickiness using the New Keynesian model of local currency pricing (theory by Betts & Devereux 2000).
- Implement the estimation method in Choi and Cook (2013).
- GMM estimation using out-of-sample forecasts of future inflation.

# Findings

- U.S. import prices are extremely sticky.
- Imports from Asia are significantly stickier than other regions.
- Differences in price stickiness can mostly explained by goods types.

#### New Keynesian Building Blocks (1)

• Calvo-type local currency pricing: the likelihood of price change = (1-v)

$$\pi_t^{LCP} = E_t \left[ \frac{(1-\nu)(1-\beta\nu)}{\nu} \left( mc_t - \mu_t \right) + \beta \cdot \pi_{t+1}^{LCP} \right]$$

- *where S:* Exchange Rate;
  - MC: Marginal Cost in Foreign Currency;
  - *IPI*: Index of Import Prices;
  - *mc:* In(MC/PPI);
  - $\mu$ : In(IPI) In(S·PPI);
  - *PPI*: Producer Prices in Foreign Currency

#### New Keynesian Building Blocks (2)

• Domestic New Keynesian Phillips Curve

$$\pi_t^{PPI} = E_t \left[ \frac{(1-\kappa)(1-\beta\kappa)}{\kappa} mc_t + \beta \pi_{t+1}^{PPI} \right]$$

• Producer Currency Pricing (PCP)

$$\pi_t^{PCP} = \pi_{t-1}^{PPI} - ds_{t-1}$$

• Combine PCP and LCP (local currency pricing)

$$\pi_t^{IPI} = \lambda \cdot \pi_t^{PCP} + (1 - \lambda) \cdot \pi_t^{LCP}$$

where  $\lambda$  : fraction of PCP pricers

#### **Estimating the Model**

$$E_{t}\left(\pi_{t}^{IPI}-\beta\pi_{t+1}^{IPI}\right)=\lambda\cdot E_{t}\left(\pi_{t}^{PCP}-\beta\cdot\pi_{t+1}^{PCP}\right)$$
$$+\left(1-\lambda\right)E_{t}\left[\frac{(1-\nu)(1-\beta\nu)\kappa}{\nu(1-\kappa)(1-\beta\kappa)}\left(\pi_{t}^{PPI}-\beta\cdot\pi_{t+1}^{PPI}\right)-\frac{(1-\nu)(1-\beta\nu)}{\nu}\mu_{t}\right].$$

• Approximate the above equation:  $\beta \sim 1$ 

$$\begin{aligned} \pi_t^{IPI} &- \pi_{t+1}^{IPI} \\ &= \alpha_0 + \alpha_1 \cdot \mu_t + \alpha_2 \cdot \left[ \pi_t^{PPI} - \pi_{t+1}^{PPI} \right] \\ &+ \alpha_3 \cdot \left[ \pi_t^{PCP} - \pi_{t+1}^{PCP} \right] + \mathcal{E}_{t+1} \end{aligned}$$

where  $\alpha_1 < 0$ ;  $\alpha_2$ ,  $\alpha_3 < 0$   $E_t[\epsilon_{t+1}] = 0$ 

### Data

- BLS Import Price Indices
- 54 U.S. trading partners
- Spot Exchange Rates, IMF IFS
- PPI Domestic Manufacturing or as close as possible.
- For country *j*,  $M_t^j \equiv \frac{S_t^j \cdot IPI_t}{PPI_t^j} \frac{1}{PPP_{BY}^j}$ , PPP relative price from ICP,  $M_t \equiv \prod_{j=1}^J (M_t^j)^{w_{t-k}^j}$ . and  $\mu_t \equiv \ln M_t$ .

### Instruments

- Expected inflation acceleration not orthogonal.
- IV (Instrumental Variables):  $\left[\pi_{t}^{PPI} \hat{\pi}_{t+1}^{PPI}\right]$ ,  $\mu_{t-1}$
- For each country *j*, construct rolling out-ofsample forecasts, depending on the availability of data,  $\hat{\pi}_{t+1}^{PPI,j}$ .
- Construct weighted average

### Import Price Stickiness (1)

- 1. European Union
- 2. Canada
- 3. Latin America
- 4. ASEAN
- 5. ANICs
- 6. Japan
- 7. China

Joint	Asian	Non Asian	
Estimation	Regions	Regions	
(A)	(B)	(C)	
0.005***	0.002***	0.016***	
(.001)	(.001)	(.003)	
0.272***	0.197***	0.465***	
(.019)	(.017)	(.047)	
0.083***	0.093***	0.072***	
(.004)	(.006)	(.008)	
474	474		
18.547	15.993		
18	15		
28.869	24.996		
0.932	0.956	0.883	
(0.005)	(0.008)	(0.008)	
0.877	0.905	0.883	
(0.009)	(0.017)	(0.008)	
	Joint Estimation (A) 0.005*** (.001) 0.272*** (.019) 0.083*** (.004) 474 18.547 18 28.869 0.932 (0.005) 0.877 (0.009)	JointAsianEstimationRegions $(A)$ $(B)$ $0.005^{***}$ $0.002^{***}$ $(.001)$ $(.001)$ $0.272^{***}$ $0.197^{***}$ $(.019)$ $(.017)$ $0.083^{***}$ $0.093^{***}$ $(.004)$ $(.006)$ $474$ $47$ $18.547$ $15.9$ $18$ $15.9$ $18$ $15.9$ $0.932$ $0.956$ $(0.005)$ $(0.008)$ $0.877$ $0.905$ $(0.009)$ $(0.017)$	

### Import Price Stickiness (2)

	Pacific				
	Rim	NICs	ASEAN	China	Japan
	(A)	(B)	(C)	(D)	(E)
$\alpha_1$	-0.003	-0.001	-0.002	-0.007	-0.005
	(.005)	(.001)	(.004)	(.008)	(.004)
$\alpha_2$	0.258***	0.096***	0.517***	0.085	-0.049
	(.059)	(.038)	(.076)	(.097)	(.057)
α <sub>3</sub>	0.061***	0.138***	0.120***	0.092***	0.057***
	(.022)	(.042)	(.027)	(.03)	(.011)
Adj. R <sup>2</sup>	0.601	0.450	0.457	0.173	0.158
Num. Obs.	34	86	34	35	87
Cragg-	15.479	17.88	11.98	9.576	42.79
Donald		Critical	Value	7.03	
V	.944	.974	.952	.920	.934
	(.036)	(.031)	(.046)	(.042)	(.021)
К	.895	.922	.934	.770	
	(.065)	(.011)	(.061)	(.132)	

### Industry-Level Regression

	<b>Independent Variable</b> $\ln(M_{2012}^{j} / M_{2002}^{j})$				
	(A)	(B)	(C)	(D)	(E)
Constant	-0.231** (.042)	-0.010 (.132)	-0.196 (.148)	-0.252** (.082)	-0.423** (.049)
Intermediate Dum	0.213* (.109)		0.192* (.111)	0.228** (.112)	
Asean+3 Share 2000		-0.331 (.242)	-0.069 (.25)		
China Share 2000				0.087 (.19)	
$\ln(\nu^j)$					-2.331*** (.711)
Average Appreciation 2000-2011					-0.019 (.026)
$\frac{N}{R^2}$	36 .100	36 .055	36 .102	36 .102	24 .633

## **Concluding Remarks**

- Consider two possibilities for U.S. import price misalignments:
  - (1) The level of misalignment is simply a function of the slow passthrough of exchange rates into import prices (Campa and Goldberg 2005; and Gust et al. 2010).
  - (2) The industries with low passthrough have not adjusted quickly to the run up in currency values of U.S. trading partners.
- Import price stickiness—rather than monetary policy per se—may explain the degree of such misalignments.
- The low passthrough of import prices from Asia is attributable to the particular composition of goods rather than any special behavior of East Asian firms.

Thank you!